

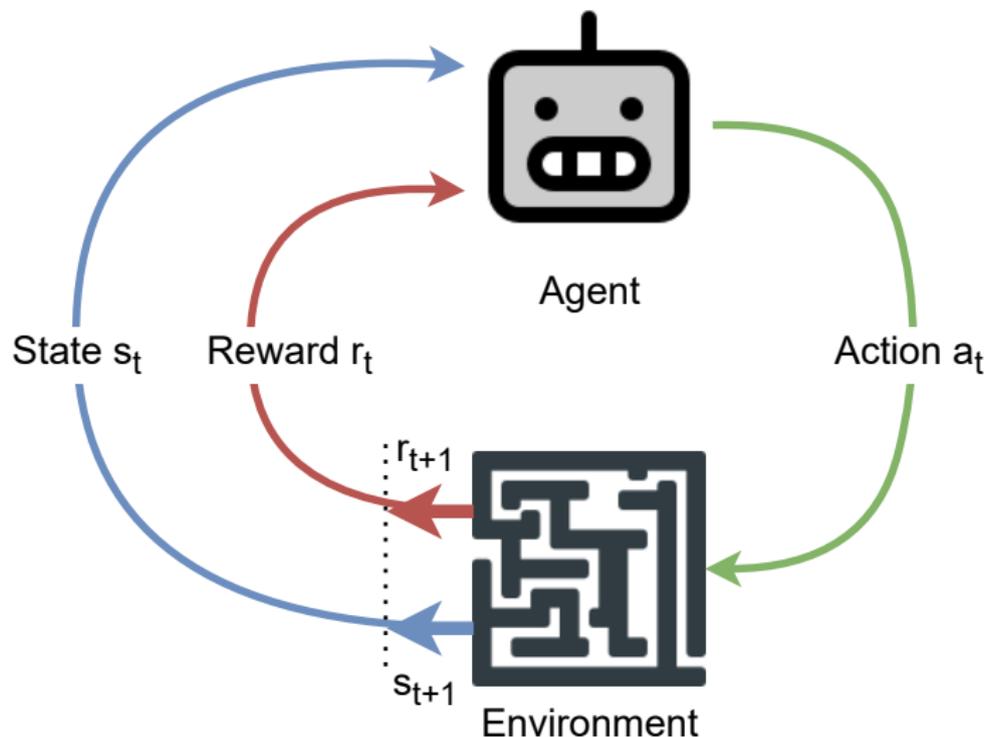
Math of Reinforcement Learning: Bayesian Approach

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Reinforcement Learning



Markov Decision Process (MDP)

Tabular, episodic MDP: H horizon, S states, A actions.

Learning in MDP: at episode t , step h

- state $s_h^t \in \mathcal{S}$;
- action $a_h^t \in \mathcal{A}$;
- next state $s_{h+1}^t \sim p_h(\cdot | s_h^t, a_h^t)$;
- reward $r_h(s_h^t, a_h^t)$ - known.

Goal: find a *policy* $\pi: \mathcal{S} \rightarrow \mathcal{A}$ that maximizes a *value function*

$$V_h^\pi(s) = \mathbb{E}_\pi \left[\sum_{h'=h}^H r_{h'}(s_{h'}, a_{h'}) \mid s_h = s \right].$$

Examples

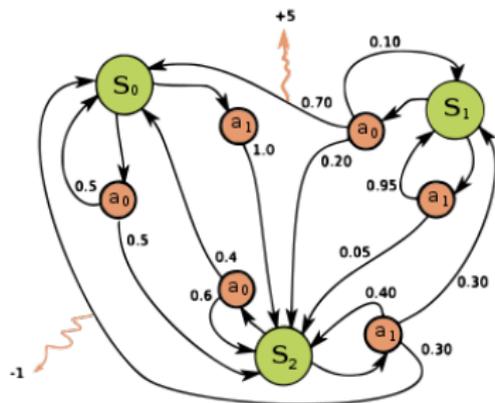


Figure: Left: MDP with $S = 3$, $A = 2$.

Right: Atari Breakout with $S = 256^{84 \cdot 84} \approx 10^{17000}$, $A = 4$.

Bellman Equations

Action-value function for policy π

$$Q_h^\pi(s, a) = \mathbb{E}_\pi \left[\sum_{h'=h}^H r_{h'}(s_{h'}, a_{h'}) \mid s_h = s, a_h = a \right].$$

Bellman equations for policy π

$$Q_h^\pi(s, a) = r_h(s, a) + p_h V_{h+1}^\pi(s, a)$$

$$V_h^\pi(s) = Q_h^\pi(s, \pi_h(s))$$

$$V_{H+1}^\pi(s) = 0$$

where $p_h f(s, a) = \sum_{s'} p_h(s'|s, a) f(s')$.

Optimal Bellman Equations

Optimal policy π^* maximizes $V_h^\pi(s)$ for all $s \in \mathcal{S}$ and $h \in [H]$.

Optimal value and action-value functions

$$V_h^*(s) = V_h^{\pi^*}(s), \quad Q_h^*(s, a) = Q_h^{\pi^*}(s, a).$$

Optimal Bellman equations

$$Q_h^*(s, a) = r_h(s, a) + p_h V_{h+1}^*(s, a)$$

$$V_h^*(s) = \max_a Q_h^*(s, a)$$

$$V_{H+1}^*(s) = 0$$

where $p_h f(s, a) = \sum_{s'} p_h(s'|s, a) f(s')$. Then $\pi_h^*(s) = \arg \max_a Q_h^*(s, a)$.

Online Reinforcement Learning Algorithm

Online algorithm: outputs a refined policy π^t after each episode $t = 1, \dots, T$.

Goal: *regret* minimization

$$\mathfrak{R}^T = \sum_{t=1}^T V_1^*(s_1^t) - V_1^{\pi^t}(s_1^t).$$

Good algorithm: sublinear regret $\mathfrak{R}^T = o(T)$.

Optimal algorithm: $\mathfrak{R}^T = \mathcal{O}(\sqrt{H^3 SAT})$ (matches the lower bound).

Exploration-Exploitation Dilemma

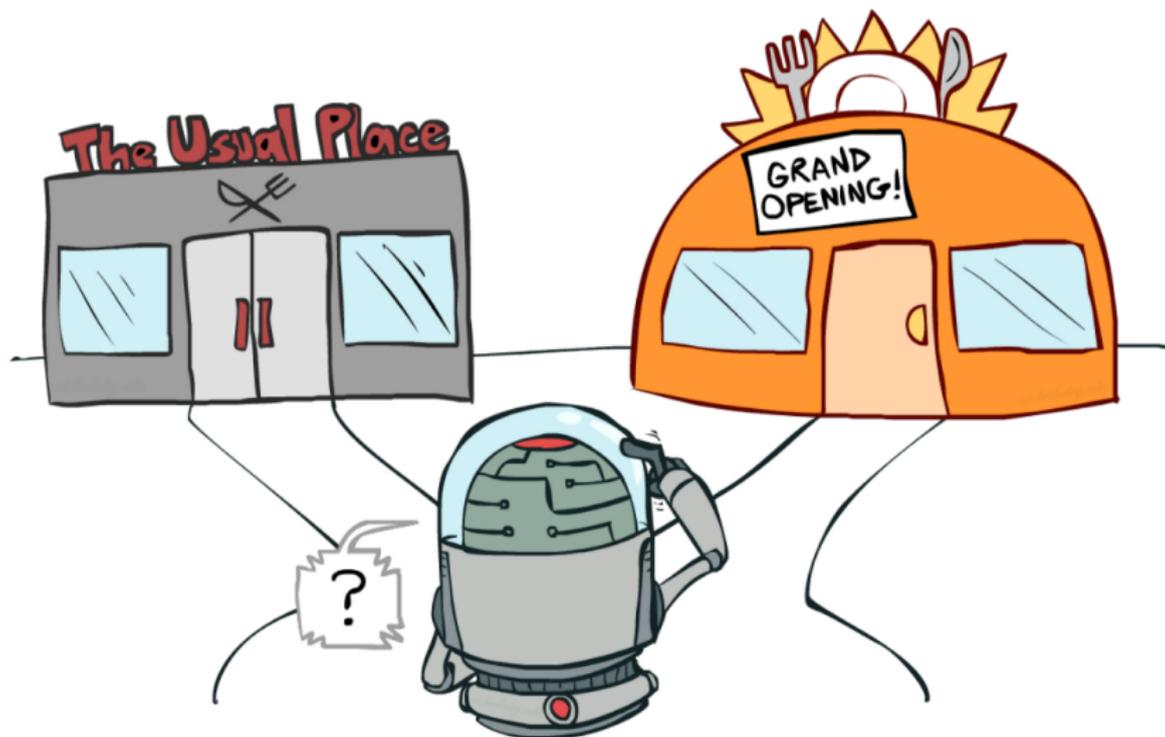


Figure: Image source: [UC Berkeley Intro to AI course](#)

Optimism in the Face of Uncertainty

Optimal Bellman Equations

$$Q_h^*(s, a) = [r_h + p_h V_{h+1}^*](s, a)$$
$$V_h^*(s) = \max_a Q_h^*(s, a)$$

- p_h - unknown!

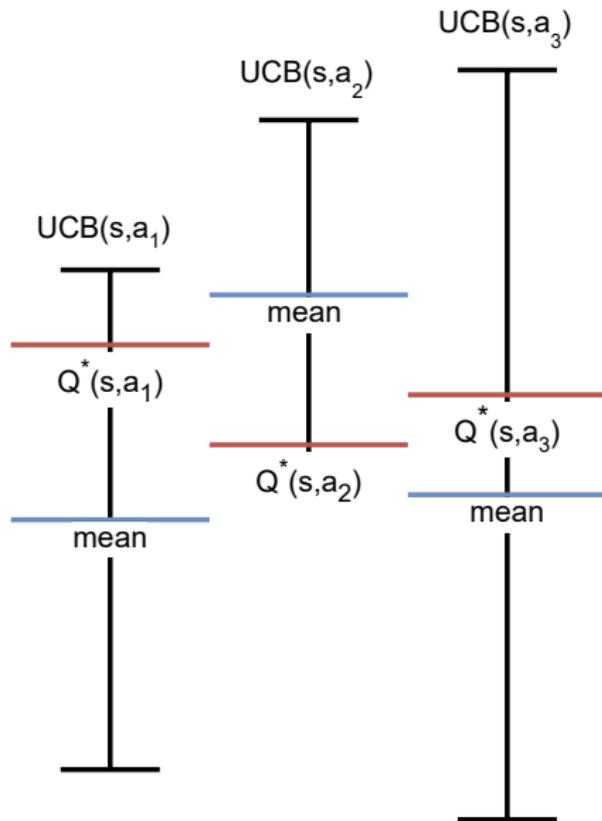
Upper confidence bound

$$\bar{Q}_h^t(s, a) = [r_h + \hat{p}_h^t \bar{V}_{h+1}^t + B_h^t](s, a)$$
$$\bar{V}_h^t(s) = \max_a \bar{Q}_h^t(s, a)$$

- \hat{p}_h^t - empirical model (mean over transitions);
- B_h^t - exploration bonus.

The most important: $\bar{Q}_h^t(s, a) \geq Q_h^*(s, a)$ with high probability.

Optimism in the Face of Uncertainty: visualization



How to Choose Bonuses: Hoeffding and Bernstein inequalities

Argument: bounded random variables concentrates near mean.

Given: X_1, \dots, X_n i.i.d. random variables, $|X_i| < b$ a.s., $\mathbb{E}[X_i] = 0$.

Theorem (Hoeffding inequality)

With probability at least $1 - \delta$ the following holds

$$\left| \frac{1}{n} \sum_{i=1}^n X_i \right| \leq \sqrt{\frac{2b^2 \log(2/\delta)}{n}}.$$

Theorem (Bernstein inequality)

With probability at least $1 - \delta$ the following holds

$$\left| \frac{1}{n} \sum_{i=1}^n X_i \right| \leq \sqrt{\frac{2\text{Var}[X_1] \log(2/\delta)}{n}} + \frac{2b \log(2/\delta)}{3n}.$$

Upper Confidence Bound Value Iteration UCBVI

[Azar et al., 2017]

Recall the setup

$$\bar{Q}_h^t(s, a) = r_h(s, a) + \underbrace{\hat{p}_h^t \bar{V}_{h+1}^t(s, a)}_{\text{upper approximation of } p_h V_{h+1}^*(s, a)} + B_h^t(s, a)$$
$$\bar{V}_h^t(s) = \max_a \bar{Q}_h^t(s, a).$$

Let $L = \log(5SAHT/\delta)$.

- UCBVI with Hoeffding bonuses

$$B_h^t(s, a) = \frac{7HL}{\sqrt{n_h^t(s, a)}}.$$

- UCBVI with Bernstein bonuses

$$B_h^t(s, a) = \sqrt{\frac{8L \text{Var}_{s' \sim \hat{p}_h^t(\cdot|s, a)}[\bar{V}_{h+1}^t(s')] }{n_h^t(s, a)}} + \frac{14HL}{3n_h^t(s, a)} + \text{correction}.$$

Near optimal in tabular setting: $\tilde{O}(\sqrt{H^3 SAT})$ regret (best up to poly-log).

UCBVI with Hoeffding bonuses: optimism proof

Lemma

For all $s, a, h, t \in \mathcal{S} \times \mathcal{A} \times [H] \times [T]$ it holds with high probability

$$\bar{Q}_h^t(s, a) \geq Q_h^*(s, a), \quad \bar{V}_h^t(s) \geq V_h^*(s).$$

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- First, by Hoeffding bound and union bound for all $s, a, h, t \in \mathcal{S} \times \mathcal{A} \times [H] \times [T]$

$$B_h^t(s, a) \geq \hat{p}_h^t V_{h+1}^*(s, a) - p_h V_{h+1}^*(s, a) \geq -B_h^t(s, a)$$

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- Next use backward induction over $h = H + 1, \dots, 1$

$$\begin{aligned} \overline{Q}_h^t(s, a) - Q_h^*(s, a) &= \hat{p}_h^t \overline{V}_{h+1}^t(s, a) + B_h^t(s, a) - p_h V_{h+1}^*(s, a) \\ &\geq \hat{p}_h^t V_{h+1}^*(s, a) + B_h^t(s, a) - p_h V_{h+1}^*(s, a) \geq 0. \end{aligned}$$

and

$$\overline{V}_h^t(s) \geq \overline{Q}_h^t(s, \pi^*(s)) \geq Q_h^*(s, \pi^*(s)) = V_h^*(s).$$

Scalability issues

Example: Go, $S \approx 10^{172}$ possible states.



Figure: Image source: Wikipedia

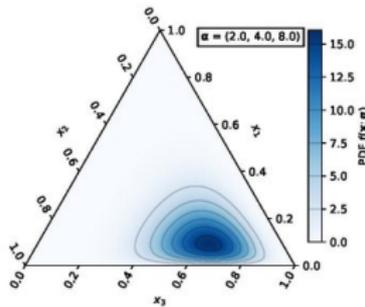
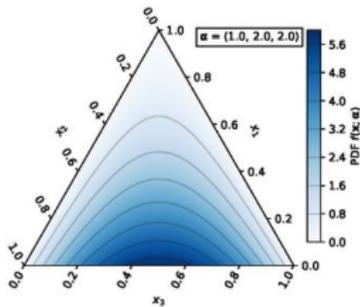
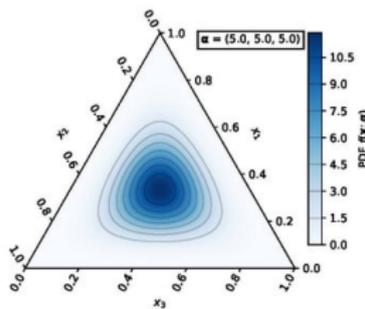
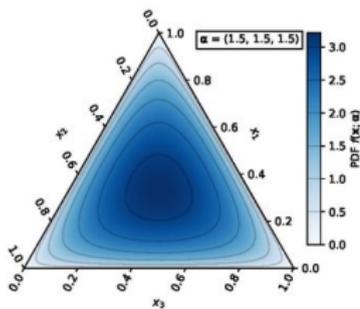
Bonus-based approach cannot be scaled: they required counters for all states.

Entering the Bayesian domain: posterior for transitions

- transitions $p_h(\cdot|s, a) \iff \text{multinomial Mult}(p_h(s'|s, a)_{s' \in \mathcal{S}})$;

Entering the Bayesian domain: posterior for transitions

- transitions $p_h(\cdot|s, a) \iff$ multinomial $\text{Mult}(p_h(s'|s, a)_{s' \in \mathcal{S}})$;
- Conjugate prior for multinomial is Dirichlet distribution: if prior $\rho_h^0(s, a)$ is $\text{Dir}(\{\bar{n}_h^0(s'|s, a)\}_{s' \in \mathcal{S}})$, then posterior $\rho_h^t(s, a)$ is $\text{Dir}(\{\bar{n}_h^0(s'|s, a) + n_h^t(s'|s, a)\}_{s' \in \mathcal{S}})$.



Preliminaries: properties of Dirichlet distribution

The Dirichlet distribution $\mathcal{Dir}(\alpha)$ for $\alpha = (\alpha_0, \dots, \alpha_m) \in \mathbb{R}_{>0}^m$ is a distribution over m -dimensional simplex $\Delta_m = \{x \in \mathbb{R}^m \mid \sum_{i=1}^m x_i \leq 1\}$

$$p(x_1, \dots, x_m) = \frac{1}{B(\alpha)} \left(1 - \sum_{i=1}^m x_i\right)^{\alpha_0 - 1} \prod_{i=1}^m x_i^{\alpha_i - 1},$$

where $B(\alpha)$ is a multivariate beta-function.

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- Representation using gamma distribution

$$(w_0, \dots, w_m) \sim \mathcal{Dir}(\alpha) \iff w_i = \frac{Y_i}{\sum_{i=0}^m Y_i}, \quad Y_i \stackrel{\text{i.i.d.}}{\sim} \Gamma(\alpha_i, 1).$$

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- Aggregation property: if $\alpha \in \mathbb{N}^m$ and $\bar{\alpha} = \sum_{i=0}^m \alpha_i$

$$\sum_{i=0}^m w_i x_i = \sum_{j=1}^{\bar{\alpha}} \hat{w}_j y_j,$$

where $w \sim \mathcal{Dir}(\alpha)$, $\hat{w} \sim \mathcal{Dir}(\mathbf{1}^{\bar{\alpha}})$, y_j are copies of x_i repeated α_i times.

Bayes-UCBVI: From Dirichlet...

Based on joint work with D.Belomenstny, E.Moulines, A.Naumov, S.Samsonov, Y.Tang, M.Valko, P.Menard. "From Dirichlet to Rubin: Optimistic Exploitation in RL without Bonuses", Oral at ICML-2022.

Idea: use directly an upper quantile over posterior distribution.

$$\bar{Q}_h^t(s, a) = r_h(s, a) + \underbrace{Q_{p \sim \rho_h^t(s, a)}}_{\text{quantile over posterior}} (p \bar{V}_{h+1}^t, \underbrace{\kappa_h^t(s, a)}_{\text{chosen quantile}})$$
$$\bar{V}_h^t(s) = \max_a \bar{Q}_h^t(s, a)$$

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- Near optimal in tabular setting: $\tilde{O}(\sqrt{H^3 SAT})$ regret.
- Scalable due to Bayesian bootstrap.

...to Rubin: Bayesian bootstrap

Given: sample $y^1, \dots, y^n \sim \mathcal{P}$.

Goal: confidence interval for $\mathbb{E}_{y \sim \mathcal{P}}[y]$.

Classical (Efron) Bootstrap

- Resample $y^{1,b}, \dots, y^{n,b}$.
- Compute mean estimate as $\frac{1}{n} \sum_{i=1}^n y^{i,b}$.
- Repeat B times.

Bayesian Bootstrap

- Sample $w^b \sim \text{Dir}(\mathbf{1}^n)$;
- Compute mean estimate as $\sum_{i=1}^n w^{b,i} y^i$;
- Repeat B times.

Then use quantiles of B mean estimates to construct a confidence interval.

Scalable implementation

- targets for Q-function estimation $y_h^n(s, a) \triangleq r_h(s, a) + \bar{V}_{h+1}^t(s_{h+1}^n)$ for $n = 1, \dots, n_h^t(s, a)$.
- prior targets $y_h^n(s, a) \triangleq r_h(s, a) + \bar{V}_h^t(s_0)$ for $n = -n_0 + 1, \dots, 0$.

By aggregation property and sample quantile approximation

$$\begin{aligned} \bar{Q}_h^t(s, a) &\triangleq r_h(s, a) + \mathbb{Q}_{p \sim \rho_h^t(s, a)}(p \bar{V}_{h+1}^t(s, a), \kappa_h^t(s, a)) \\ &= \mathbb{Q}_{w \sim \text{Dir}(\mathbf{1}^{\bar{n}_h^t(s, a)})} \left(\sum_{n=-n_0+1}^{n_h^t(s, a)} w_n y_h^n(s, a), \kappa_h^t(s, a) \right) \\ &\approx \underbrace{\mathbb{Q}_{b \sim \text{Unif}([B])} \left(\sum_{n=-n_0+1}^{n_h^t(s, a)} w_h^{n, b}(s, a) y_h^n(s, a), \kappa_h^t(s, a) \right)}_{\text{upper confidence bound by Bayesian bootstrap}}. \end{aligned}$$

Deep RL extension: Bayes-UCBDQN

Recall

$$\bar{Q}_h^t(s, a) \approx \mathbb{Q}_{b \sim \mathcal{U}(\text{nif}([B]))} \left(\bar{Q}_h^{t,b}(s, a), \kappa_h^t(s, a) \right)$$
$$\text{where } \bar{Q}_h^{t,b}(s, a) \triangleq \sum_{n=-n_0+1}^{n_h^t(s,a)} w_h^{n,b}(s, a) y_h^n(s, a).$$

Uniform Dirichlet distribution = exponential ($\Gamma(1, 1)$) with normalization

$$\bar{Q}_h^{t,b}(s, a) = \arg \min_x \sum_{n=-n_0+1}^{n_h^t(s,a)} z_h^{n,b}(s, a) (x - y_h^n(s, a))^2$$

where $z_h^{n,b}(s, a) \sim \mathcal{E}(1)$ i.i.d. .

Deep RL:

- sample minibatch of targets;
- update parameters by the gradient of weighted linear regression.

Experimental results

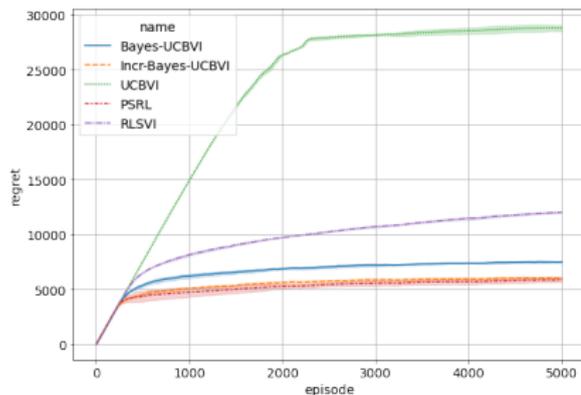


Figure: Left: Regret of Bayes-UCBVI and Incr-Bayes-UCBVI compared to baselines on grid-world with 5 rooms of size 5×5 . Right: deep RL algorithms with median human normalized scores across Atari-57 games.

Back to theory: optimistic prior

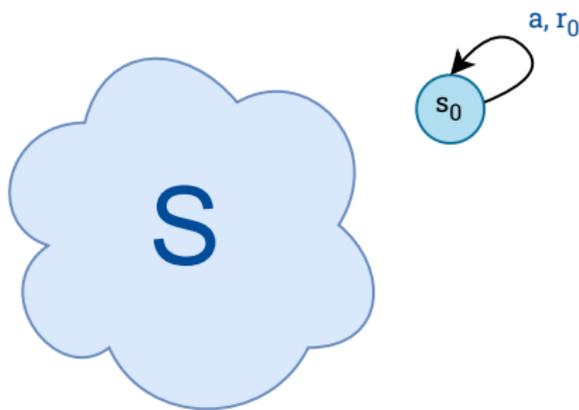


Figure: Extended state space by a fake state s_0 , $r_0 > 1$.

Goal: encourage initial exploration.

- *Tabular:* prior $\rho_h^0(s'|s, a) = \text{Dir}(\{n_0\}_{s'=s_0} \cup \{0\}_{s' \in \mathcal{S}})$.
- *Deep RL:* Add n_0 prior transitions to s_0 ;

Theoretical analysis

Let us fix $\delta \in (0, 1)$, $r_0 \triangleq 2$, $n_0 \triangleq \mathcal{O}(\log(T))$, and the quantile function

$$\kappa_h^t(s, a) \triangleq 1 - \frac{C_\kappa \delta}{\underbrace{SAH[2n_h^t(s, a) + 1]^3 [\bar{n}_h^t(s, a)]^{3/2}}_{\text{polynomial in parameters}}}.$$

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Theorem (Regret bound)

For Bayes-UCBVI, with probability at least $1 - \delta$,

$$\mathfrak{R}^T = \mathcal{O}\left(\sqrt{H^3 SATL} + H^3 S^2 AL^2\right),$$

where $L \triangleq \mathcal{O}(\log(HSAT/\delta))$.

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where $L \triangleq \mathcal{O}(\log(HSAT/\delta))$.

Matches the lower bound $\Omega(\sqrt{H^3 SAT})$ up to poly-log terms.

Sketch of proof

The heart of the analysis is a novel anti-concentration inequality.

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Theorem (Dirichlet boundary crossing, Informal)

For any $\alpha = (\alpha_0, \alpha_1, \dots, \alpha_m) \in \mathbb{N}^{m+1}$ define $\bar{p} \in \Delta_m$ with $\bar{p}(\ell) = \alpha_\ell / \bar{\alpha}$, $\ell = 0, \dots, m$, where $\bar{\alpha} = \sum_{j=0}^m \alpha_j$. Under technical assumptions, for $f: \{0, \dots, m\} \rightarrow [0, b_0]$ and $\mu \in (\bar{p}f, b_0)$

$$\frac{\exp(-\bar{\alpha} \mathcal{K}_{inf}(\bar{p}, \mu, f))}{\bar{\alpha}^{3/2}} \leq \mathbb{P}_{w \sim \text{Dir}(\alpha)}[wf \geq \mu] \leq \exp(-\bar{\alpha} \mathcal{K}_{inf}(\bar{p}, \mu, f)),$$

where $\mathcal{K}_{inf}(p, u, f)$ is given by

$$\mathcal{K}_{inf}(p, u, f) \triangleq \max_{\lambda \in [0, 1]} \mathbb{E}_{X \sim p} \left[\log \left(1 - \lambda \frac{f(X) - u}{b_0 - u} \right) \right].$$

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- Lower bound is an essential part for optimism;
- Upper bound is important for the reduction to UCBVI.

Takeaways

- Optimism in the face of uncertainty principle as a solution to exploration-exploitation dilemma;
- Bayesian perspective gives more possibility to scale up algorithms;
- Reinforcement learning is full of mathematical questions and fun!

Thank you!

Bibliography I

-  Azar, M. G., Osband, I., and Munos, R. (2017). [Minimax regret bounds for reinforcement learning](#).
In International Conference on Machine Learning.